

# Universal Off-Policy Evaluation



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Games (**Expected** performance)

Mean







You migh	nt also like

Games (**Expected** performance) Automated healthcare (Mitigate **risk**)

Mechanical control (Mitigate uncertainty) Online recommendation (Robust to **noisy** data-collection)

Mean

VaR, CVaR

Variance, Entropy

Median, Inter-quantiles







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Markovian

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**Partial Observation** 

Variance, Entropy

Non-Markovian

Median, Inter-quantiles

**Non-stationary** 

Real-world problems are often high-stakes. Evaluate a policy's performance *before* deployment (**off-policy**).







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**Given**: Trajectories collected using one or multiple *past* (behavior) policies.

## Goal

**Given**: Trajectories collected using one or multiple *past* (behavior) policies.

**Aim**: Evaluate and bound the desired performance metric (mean, variance, CVaR, etc.) of the return distribution under a *new* policy, using the given trajectories.



## Goal

#### Assumptions

- Any outcome under the evaluation policy is possible under the behavior policy (support assumption)
- Knowledge of action probabilities under the behavior policy



## **Prior work**

- Model-based
  - Additional requirement for estimating reward distribution for each state-action pair
  - Hard to estimate accurate models in **non-tabular** settings
- Typical IS based estimators
  - Only corrects for the mean
- Distributional RL
  - On-policy

## **A Universal Evaluation Procedure**

- Off-policy
  - Model-free
- Different **performance metrics** (Estimation + High-confidence bounds)
  - Mean,
  - VaR, CVaR,
  - Variance, Entropy,
  - Median, Inter-quantiles
  - **Etc**.
- Different domain settings
  - Markovian, Non-Markovian
  - Fully observable, Partially-observable
  - Smoothly non-stationary, discrete distribution shifts

#### **Core Idea**

- If we have an **estimator for the CDF** then we can obtain an estimator for any of its parameters

$$F_{\pi}(\nu) \coloneqq \Pr\left(G_{\pi} \leq \nu\right), \qquad \forall \nu \in \mathbb{R}$$

- Bounds for the CDF can directly be used to obtain bounds on its parameters

$$\Pr(orall 
u \in \mathbb{R}, F_\pi(
u) \in \mathcal{F}(
u)) \geq 1-\delta$$

#### **Intuition for CDF Estimator**



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#### **CDF Estimator**

**Theorem 1.** Under Assumption 1,  $\hat{F}_n$  is an unbiased and uniformly consistent estimator of  $F_{\pi}$ . That is,

$$\forall \nu \in \mathbb{R}, \quad \mathbb{E}_{\mathcal{D}} \Big[ \hat{F}_n(\nu) \Big] = F_{\pi}(\nu),$$
$$\sup_{\nu \in \mathbb{R}} \left| \hat{F}_n(\nu) - F_{\pi}(\nu) \right| \xrightarrow{a.s.} 0.$$

#### **Estimates for Different Parameters**

$$\hat{F}_n^{-1}(\alpha) \coloneqq \min \Big\{ g \in (G_{(i)})_{i=1}^n \Big| \hat{F}_n(g) \ge \alpha \Big\},$$
  
CDF Inverse

$$d\hat{F}_n(G_{(i)}) \coloneqq \hat{F}_n(G_{(i)}) - \hat{F}_n(G_{(i-1)}),$$

$$\mathsf{PDF}$$

#### **Estimates for Different Parameters**

$$\hat{F}_n^{-1}(\alpha) \coloneqq \min\left\{g \in (G_{(i)})_{i=1}^n \middle| \hat{F}_n(g) \ge \alpha\right\}, \qquad \mathrm{d}\hat{F}_n(G_{(i)}) \coloneqq \hat{F}_n(G_{(i)}) - \hat{F}_n(G_{(i-1)}),$$
CDF Inverse PDF

$$\begin{split} \mu_{\pi}(\hat{F}_{n}) &\coloneqq \sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)})G_{(i)}, & Q_{\pi}^{\alpha}(\hat{F}_{n}) \coloneqq \hat{F}_{n}^{-1}(\alpha), \\ \mathrm{IQR}_{\pi}^{\alpha_{1},\alpha_{2}}(\hat{F}_{n}) &\coloneqq Q_{\pi}^{\alpha_{2}}(\hat{F}_{n}) - Q_{\pi}^{\alpha_{1}}(\hat{F}_{n}), \\ \sigma_{\pi}^{2}(\hat{F}_{n}) &\coloneqq \sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)}) \left(G_{(i)} - \mu_{\pi}(\hat{F}_{n})\right)^{2}, & \mathrm{IQR}_{\pi}^{\alpha_{1},\alpha_{2}}(\hat{F}_{n}) \coloneqq Q_{\pi}^{\alpha_{2}}(\hat{F}_{n}) - Q_{\pi}^{\alpha_{1}}(\hat{F}_{n}), \\ \mathcal{H}_{\pi}(\hat{F}_{n}) &\coloneqq -\sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)}) \log \mathrm{d}\hat{F}_{n}(G_{(i)}). & \mathrm{CVaR}_{\pi}^{\alpha}(\hat{F}_{n}) \coloneqq \frac{1}{\alpha} \sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)})G_{(i)}\mathbb{1}_{\{G_{(i)} \leq Q_{\pi}^{\alpha}(\hat{F}_{n})\}}. \end{split}$$

#### **Estimates for Different Parameters**

$$\hat{F}_n^{-1}(\alpha) \coloneqq \min \Big\{ g \in (G_{(i)})_{i=1}^n \Big| \hat{F}_n(g) \ge \alpha \Big\}, \qquad \mathrm{d}\hat{F}_n(G_{(i)}) \coloneqq \hat{F}_n(G_{(i)}) - \hat{F}_n(G_{(i-1)}),$$

Mean estimate *exactly* equal to the common (trajectory-based) IS estimate.

$$\begin{split} \mu_{\pi}(\hat{F}_{n}) &\coloneqq \sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)})G_{(i)}, & Q_{\pi}^{\alpha}(\hat{F}_{n}) \coloneqq \hat{F}_{n}^{-1}(\alpha), \\ \mathrm{IQR}_{\pi}^{\alpha_{1},\alpha_{2}}(\hat{F}_{n}) &\coloneqq Q_{\pi}^{\alpha_{2}}(\hat{F}_{n}) - Q_{\pi}^{\alpha_{1}}(\hat{F}_{n}), \\ \sigma_{\pi}^{2}(\hat{F}_{n}) &\coloneqq \sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)}) \Big(G_{(i)} - \mu_{\pi}(\hat{F}_{n})\Big)^{2}, & \mathrm{CVaR}_{\pi}^{\alpha}(\hat{F}_{n}) \coloneqq \frac{1}{\alpha} \sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)})G_{(i)}\mathbb{1}_{\{G_{(i)} \leq Q_{\pi}^{\alpha}(\hat{F}_{n})\}} \\ \mathcal{H}_{\pi}(\hat{F}_{n}) &\coloneqq -\sum_{i=1}^{n} \mathrm{d}\hat{F}_{n}(G_{(i)}) \log \mathrm{d}\hat{F}_{n}(G_{(i)}). \end{split}$$

## **Bounds for Different Parameters**

- Mean
- Quantile
- CVaR
- Inter-quantile
- Entropy
- Variance
- ....

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- ....

- Estimates for different parameters might be **biased**
- Importance sampling results in high variance
- Need to obtain high-confidence bounds with guaranteed coverage for reliability.

## **Bounds**



## **Bounds**

$${\hat F}_n(\kappa):=rac{1}{n}\sum_{i=1}^n
ho_i(1_{\{G_i\leq\kappa\}})$$
 ,

#### **Mean Estimation!**

Let 
$$X \coloneqq \rho(1_{\{G \le \kappa\}})$$
.  
 $\mathbb{E}_{\mathcal{D}}[X] = F_{\pi}(\kappa)$ .







#### **Bounds**

**Theorem 3.** Under Assumption 1, for any  $\delta \in (0, 1]$ , if  $\sum_{i=1}^{K} \delta_i \leq \delta$ , then the confidence band defined by  $F_-$  and  $F_+$  provides guaranteed coverage for  $F_{\pi}$ . That is,  $\Pr\left(\forall \nu, F_-(\nu) \leq F_{\pi}(\nu) \leq F_+(\nu)\right) \geq 1 - \delta.$ 















## Bootstrap

**Algorithm 1:** Bootstrap Bounds for  $\psi(F_{\pi})$ 

- 1 Input: Dataset  $\mathcal{D}$ , Confidence level  $1 \delta$
- <sup>2</sup> Bootstrap *B* datasets  $\{\mathcal{D}_i^*\}_{i=1}^B$  and create  $\{\bar{F}_{n,i}^*\}_{i=1}^B$
- 3 Bootstrap estimates  $\{\psi(\bar{F}_{n,i}^*)\}_{i=1}^B$  using  $\{\bar{F}_{n,i}^*\}_{i=1}^B$ .
- 4 Compute  $\{\psi_{-}, \psi_{+}\}$  using BCa $(\{\psi(\bar{F}_{n,i}^{*})\}_{i=1}^{B}, \delta)$  [1]
- 5 Return  $\{\psi_-,\psi_+\}$ 
  - Approximate
  - Significantly Tighter

[1] DiCiccio, Thomas J., and Bradley Efron. "Bootstrap confidence intervals." Statistical science 11.3 (1996)



#### 30k samples 30 Trials



[1] Thomas, Philip, Georgios Theocharous, and Mohammad Ghavamzadeh. "High-confidence off-policy evaluation." AAAI 2015.

[2] Chandak, Yash, Shiv Shankar, and Philip S. Thomas. "High-Confidence Off-Policy (or Counterfactual) Variance Estimation." AAAI 2021.





## **Extensions**

- Weighted IS based UnO for variance reduction\*
- UnO for partially observable MDPs\*
- UnO for discrete distributional shifts\*
- UnO for smooth non-stationarities\*
- Parallel work at NeurIPS'21 by Audrey et al. [1] provides uniform convergence rates for off-policy CDF and Lipschitz risk functionals.

\*see our paper for more details.

[1] Huang, Audrey, Liu Leqi, Zachary C. Lipton, and Kamyar Azizzadenesheli. "Off-Policy Risk Assessment in Contextual Bandits." NeurIPS 2021.

