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Asymptotic Guarantees for Learning Generative Models with the Sliced-Wasserstein Distance

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Minimum Distance Estimation

$$\hat{\theta}_n = \operatorname{argmin}_{\theta} D(\hat{\mu}_n; \theta)$$

D : distance between distributions

$\hat{\mu}_n$: empirical distribution of data points Y_1, \dots, Y_n i.i.d from θ

θ : distribution parametrized by θ

Minimum Distance Estimation

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D : distance between distributions

Example: **Generative Modeling**

$\hat{\mu}_n$: empirical distribution of **data points** Y_1, \dots, Y_n i.i.d from ?

θ : distribution parametrized by θ

Minimum Expected Distance Estimation

Directly optimizing $\hat{\theta}_{n;m}$ is often not possible (e.g. GANs)

$$\hat{\theta}_{n;m} = \operatorname{argmin}_{\theta} E [D(\hat{\mu}_n; \hat{\mu}_{\theta;m}) | Y_{1:n}]$$

$\hat{\mu}_{\theta;m}$: empirical distribution of a sample $Z_1; \dots; Z_m$ i.i.d. from

Minimum Wasserstein Estimation

Choose $D = W_p$ (Wasserstein distance of order $p \geq 1$)

- X Robust and increasingly popular estimators: Wasserstein GAN [1], Wasserstein auto-encoders [2]
- X Asymptotic guarantees [3]

[1] Arjovsky et al., 2017 [2] Tolstikhin et al., 2018 [3] Bernton et al., 2019

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W_p : expensive + curse of dimensionality

Central limit theorem in [3] valid in 1D

Sliced-Wasserstein distance

In 1D, W_p has an analytical form) Motivates a practical alternative:

$$SW_p^p(\mu; \nu) = \int_{S^{d-1}} W_p^p(u_j^?; u_j^?) d(u)$$

Minimum Sliced-Wasserstein Estimation

$$\hat{\mu}_n = \operatorname{argmin}_{\mu} \mathbb{E} [\operatorname{SW}_p(\hat{\mu}_n; \mu)]$$
$$\hat{\mu}_{n;m} = \operatorname{argmin}_{\mu} \mathbb{E} [\operatorname{SW}_p(\hat{\mu}_n; \mu) \mid Y_{1:n}]$$

Successful in generative modeling applications (e.g., SW-GAN, Deshpande et al., 2018)

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Our contributions:


Convergence in SW_p weak convergence of probability measures

Existence and consistency of $\hat{\mu}_n, \hat{\mu}_{n;m}$

Central limit theorem for $\hat{\mu}_n$: \sqrt{n} convergence rate for any dimension

Thank you!

Our Poster: **East Exhibition Hall B + C #226**




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Minimum Distance Estimation

Observations $Y_{1:n} = (Y_1, \dots, Y_n)$, $Y_i \in \mathbb{R}^d$, i.i.d. from $\mu \in \mathcal{P}(\mathbb{R}^d)$, with $P(Y) = \mu$, set of probability measures on \mathbb{R}^d .

A family of distributions on \mathbb{R}^d parameterized by $\theta \in \mathbb{R}^p$: $\mu_\theta \in \mathcal{P}(\mathbb{R}^d)$; $\theta \in \Theta$.

Purely generative models: We can generate $2N$ i.i.d. samples from μ_θ , but the likelihood is intractable. μ_n is the empirical distribution.

Given $Y_{1:n}$, its empirical distribution μ_n , and a distance D on $\mathcal{P}(\mathbb{R}^d)$, we perform Minimum Distance Estimation (MDE):

$$\hat{\theta}_n = \arg \min_{\theta \in \Theta} D(\mu_n, \mu_\theta) \quad (1)$$

or Minimum Expected Distance Estimation (MEDE):

$$\hat{\theta}_n = \arg \min_{\theta \in \Theta} E[D(\mu_n, \mu_\theta) | Y_{1:n}] \quad (2)$$

Theoretical Results

The convergence MDE implies the weak convergence $\text{MP}(\mathbb{R}^d)$.

Key assumptions.

Continuity: For any $(\theta_n)_{n \geq 1}$ such that $\lim_{n \rightarrow \infty} \theta_n = \theta$, $\mu_{\theta_n} \rightarrow \mu_\theta$ converges weakly if $\int \phi(x) d\mu_{\theta_n}(x) \rightarrow \int \phi(x) d\mu_\theta(x)$.

A1: $\int \phi(x) d\mu_{\theta_n}(x) \rightarrow \int \phi(x) d\mu_\theta(x)$.

A2: $\lim_{n \rightarrow \infty} \int \phi(x) d\mu_{\theta_n}(x) = \int \phi(x) d\mu_\theta(x)$.

Data-generating process:

A3: $\lim_{n \rightarrow \infty} \int \phi(x) d\mu_{\theta_n}(x) = \int \phi(x) d\mu_\theta(x)$, P -almost surely.

Bounded rate: For some $\gamma > 0$,


A4: $\int \phi(x) d\mu_{\theta_n}(x) - \int \phi(x) d\mu_\theta(x) \leq C \|\theta_n - \theta\|^\gamma$, is bounded.

A5: $\int \phi(x) d\mu_{\theta_n}(x) - \int \phi(x) d\mu_\theta(x) \leq C \|\theta_n - \theta\|^\gamma$, is bounded almost surely.


Numerical Experiments

Multivariate Gaussians.


$\mu = N(m, \Sigma)$; $m \in \mathbb{R}^d$; $\Sigma > 0$, $d \geq 1$, and $(m, \Sigma) \in \mathcal{D}$.




MSE vs. n



MSE vs. n = m




MSE vs. n = 2000 vs. n




Multivariate elliptically contoured stable distributions.


$\mu = E S_\alpha(\sigma, m)$ with $\alpha \in (1, 2]$, and $m, \sigma \in \mathbb{R}^d$.




Comparison Wasserstein and SW



MSE vs. n




MSE vs. n = 100



High-dimensional real data.

We train the Sliced-Wasserstein Generator [2] (based on MESWE).

We plot the mean-squared error between the mean-squared error obtained for (n, m) from $(1, 1)$ to $(10, 0.005)$ and for $(n, m) = (50000, 200)$.



Optimal Transport (OT) Metrics

For $p \geq 1$, $P_p(\mathbb{R}^d)$: set of probability measures on \mathbb{R}^d with finite p -th moment. Let $\mu \in P_p(\mathbb{R}^d)$.

Wasserstein distance (W_p) . Computationally expensive, except in 1D (\mathbb{R}) analytical form.

Sliced-Wasserstein (SW) distance

\mathbb{S}^d : d -dimensional unit sphere, uniform distribution on \mathbb{S}^d .

Practical metric based on projections:

$$\text{SW}_p(\mu, \nu) = \int_{\mathbb{S}^d} W_p(\mu|_u, \nu|_u) du$$

Existence and consistency of MSWE

Assume A1, A3, A4. Then, there exist $(\theta_n)_{n \geq 1}$ such that, for all $\theta \in \Theta$,

$$\int \phi(x) d\mu_{\theta_n}(x) \rightarrow \int \phi(x) d\mu_\theta(x)$$

$$\limsup_{n \rightarrow \infty} \int \phi(x) d\mu_{\theta_n}(x) \leq \int \phi(x) d\mu_\theta(x)$$

Besides, for all $\theta \in \Theta$, there exist $(\theta_n)_{n \geq 1}$ such that, for all $n \geq 1$, $\theta_n \in \Theta$ and $\int \phi(x) d\mu_{\theta_n}(x) \rightarrow \int \phi(x) d\mu_\theta(x)$ is non-empty.

Combining MDE and OT

Minimum Wasserstein estimators, defined in (1) and (2) with $D = W_p$, have asymptotic guarantees [1] but are not practical.

With $D = \text{SW}_p$ in (1) and (2), we get the minimum (expected) SW estimators $\text{MDE}(\text{SW}_p)$ of order p .

Recent studies show the empirical success of SW-based estimators on generative modeling but lack of theoretical guarantees.

) We investigate the asymptotic properties of these estimators.

Guarantees for MSWE

Existence and consistency (with A1 to A6), convergence to MSWE as $m \rightarrow \infty$ (A1, A2, A5).

Central limit theorem for MSWE with $p = 1$

Consider A1, A3, A4, $\theta \in \Theta$ with $\theta \in \Theta$ self-separated and $H = \int \phi(x) d\mu_\theta(x)$ with $D(\mu_\theta, \mu_\theta) > 0$, with $\mathbb{P}(\theta_n \in \Theta) \rightarrow 1$ as $n \rightarrow \infty$, where θ_n and θ contain the CDFs of the projected μ_n and μ_θ .

$D(\mu_n, \mu_\theta)$: the derivative of $F(\theta)$ in θ .

Then, $\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, \Sigma)$ with $\Sigma = \frac{1}{D(\mu_\theta, \mu_\theta)^2} \text{Cov}(\phi(X))$.

) Convergence rate of $\hat{\theta}_n$ independent of the dimension

Main References

[1] E. Bertoin, P. E. Jacot, M. Gerber, C. P. Robert. On generator estimation with the Wasserstein distance. Information and Inference: A Journal of the IMA, Jan 2018.

[2] I. Dierker, Z. Zhang, A. G. Schwing. Generative modeling using the sliced Wasserstein distance. CVPR 2018.